

A Study on Analytical Properties of Bayesian Experimental Design Model based on an Orthonormal System

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Motivation and Objective

In signal processing, the model is generally based on orthonormal systems. In such models

- All parameters are independent.
- The fast Fourier Transform (FFT) can calculate the parameters.
- Parameters are complex numbers.

In some previous experimental design models

- Parameters are not necessarily independent.
- Parameters are real numbers.

Motivation

Can we use the idea of orthonormal systems to calculate the posterior variance in Bayesian experimental designs?

Objective

We focus on a subclass of designs, which is limited by orthonormal systems. Then, it's to show we can get the posterior variance directly.

Orthonormal Systems: Example 1

F_1, F_2, \dots, F_n : Factors

$x_i \in \{0, 1\}$: the level of F_i

$\mathbf{x} = (x_1, x_2, \dots, x_n) \in \{0, 1\}^n$: level combination

$\{0, 1\}^n$: the set of all sequences of 0, 1 that is n long

Basis functions over the Boolean domain

For each $\mathbf{a} = (a_1, a_2, \dots, a_n) \in \{0, 1\}^n$, define the basis function

$$\chi_{\mathbf{a}}(\mathbf{x}) = (-1)^{\mathbf{a} \cdot \mathbf{x}},$$

where $\mathbf{a} \cdot \mathbf{x} = a_1x_1 + a_2x_2 + \dots + a_nx_n$ is to be performed modulo 2.

Orthonormal Systems

The basis functions form orthonormal systems, that is,

$$\frac{1}{2^n} \sum_{\mathbf{x} \in \{0, 1\}^n} \chi_{\mathbf{a}}(\mathbf{x}) \chi_{\mathbf{b}}(\mathbf{x}) = \begin{cases} 1, & \mathbf{a} = \mathbf{b}, \\ 0, & \mathbf{a} \neq \mathbf{b}. \end{cases}$$

Orthonormal Systems: Example 2

F_1, F_2, \dots, F_n : Factors

$x_i \in \{0, 1, 2\}$: the level of F_i

$\mathbf{x} = (x_1, x_2, \dots, x_n) \in \{0, 1, 2\}^n$: level combination

$\{0, 1, 2\}^n$: the set of all sequences of 0, 1, 2 that is n long

Basis functions over $\{0, 1, 2\}^n$ domain ($GF(3)^n$ domain)

For each $\mathbf{a} = (a_1, a_2, \dots, a_n) \in \{0, 1, 2\}^n$, define the basis function

$$\chi_{\mathbf{a}}(\mathbf{x}) = e^{i2\pi\mathbf{a}\cdot\mathbf{x}/3},$$

where $\mathbf{a} \cdot \mathbf{x} = a_1x_1 + a_2x_2 + \dots + a_nx_n$ is to be performed over $GF(3)$.

Orthonormal Systems

The basis functions form orthonormal systems, that is,

$$\frac{1}{3^n} \sum_{\mathbf{x} \in \{0,1,2\}^n} \chi_{\mathbf{a}}(\mathbf{x})\chi_{\mathbf{b}}^*(\mathbf{x}) = \begin{cases} 1, & \mathbf{a} = \mathbf{b}, \\ 0, & \mathbf{a} \neq \mathbf{b}, \end{cases}$$

where $\chi_{\mathbf{b}}^*(\mathbf{x})$ is the complex conjugate of $\chi_{\mathbf{b}}(\mathbf{x})$.

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Example: Model of Two-factor Experiments

F_1, F_2 : the factors

$x_i \in \{0, 1, 2\}$: the level of F_i

$\mathbf{x} = (x_1, x_2) \in \{0, 1, 2\}^2$: level combination

$t(\mathbf{x})$: the response of the experiment with level combination \mathbf{x}

Previous model

$$t(\mathbf{x}) = \mu + \alpha_1(x_1) + \alpha_2(x_2) + \beta_{1,2}(x_1, x_2) + \epsilon,$$

where

μ : the general mean

$\alpha_1(x_1)$: the effect of the x_1 level of F_1

$\alpha_2(x_2)$: the effect of the x_2 level of F_2

$\beta_{1,2}(x_1, x_2)$: the interaction of the x_1 level of F_1 and the x_2 level of F_2

ϵ : a zero-mean Gaussian random variable with variance σ^2

Parameters on Previous Model

Then, all parameters are given as follows:

$\mu, \alpha_1(0), \alpha_1(1), \alpha_1(2), \alpha_2(0), \alpha_2(1), \alpha_2(2), \beta_{1,2}(0,0), \beta_{1,2}(0,1), \beta_{1,2}(0,2),$
 $\beta_{1,2}(1,0), \beta_{1,2}(1,1), \beta_{1,2}(1,2), \beta_{1,2}(2,0), \beta_{1,2}(2,1), \beta_{1,2}(2,2).$

independent parameters vector $\mathbf{u} \in \mathbb{R}^9$

Under the constraints on parameters, let \mathbf{u} denote the independent parameters vector.

$$\mathbf{u} = \begin{bmatrix} \mu \\ \alpha_1(0) \\ \alpha_1(1) \\ \alpha_2(0) \\ \alpha_2(1) \\ \beta_{1,2}(0,0) \\ \beta_{1,2}(0,1) \\ \beta_{1,2}(1,0) \\ \beta_{1,2}(1,1) \end{bmatrix},$$

Model based on orthonormal systems (Ukita et al. 2010)

$$t(\mathbf{x}) = \sum_{\mathbf{a} \in \{0,1,2\}^2} f_{\mathbf{a}} \chi_{\mathbf{a}}(\mathbf{x}) + \epsilon,$$

where

$f_{\mathbf{a}} : f_{\mathbf{a}} \in \mathbb{C}^9$

$\chi_{\mathbf{a}}(\mathbf{x}) = e^{i2\pi \mathbf{a} \cdot \mathbf{x} / 3}$, and

ϵ : a zero-mean Gaussian random variable with variance σ^2

Then, the basis functions $\{\chi_{\mathbf{a}} | \mathbf{a} \in \{0, 1, 2\}^2\}$ form orthonormal systems.

Parameters (Model based on orthonormal systems)

independent parameters vector $\mathbf{w} \in \mathbb{C}^9$ (Fourier coefficients vector)

$$\mathbf{w} = \begin{bmatrix} f_{00} \\ f_{10} \\ f_{20} \\ f_{01} \\ f_{02} \\ f_{11} \\ f_{12} \\ f_{21} \\ f_{22} \end{bmatrix} .$$

no constraints on parameters \rightarrow all parameters are independent.

Relation between \mathbf{u} and \mathbf{w}

There is a 9×9 matrix \mathbf{M} that satisfies

$$\mathbf{u} = \mathbf{M}\mathbf{w},$$

and

$$\mathbf{M} = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & \omega_3 & \omega_3^2 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & \omega_3 & \omega_3^2 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 1 & 1 & 1 \\ 0 & 0 & 0 & 0 & 0 & \omega_3 & \omega_3^2 & \omega_3 & \omega_3^2 \\ 0 & 0 & 0 & 0 & 0 & \omega_3 & \omega_3 & \omega_3^2 & \omega_3^2 \\ 0 & 0 & 0 & 0 & 0 & \omega_3^2 & 1 & 1 & \omega_3 \end{bmatrix}.$$

where $\omega_3 = e^{2\pi i/3}$. \mathbf{M}^{-1} also exists and $\mathbf{M}^{-1}\mathbf{u} = \mathbf{w}$.

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Likelihood Function

- a data set of inputs $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$ with corresponding target values $t(\mathbf{x}_1), \dots, t(\mathbf{x}_N)$
- the variables $\{t(\mathbf{x}_1), \dots, t(\mathbf{x}_N)\}$ be a column vector denoted by \mathbf{t}
- K : the number of parameters

Likelihood Function

$$p(\mathbf{t}|\mathbf{X}, \mathbf{u}, \sigma^2) = \mathcal{N}(\mathbf{t}|\Phi\mathbf{M}^{-1}\mathbf{u}, \sigma^2\mathbf{I}).$$

where

$$\Phi = \begin{bmatrix} \mathcal{X}_{a_1}(\mathbf{x}_1) & \mathcal{X}_{a_2}(\mathbf{x}_1) & \dots & \mathcal{X}_{a_K}(\mathbf{x}_1) \\ \mathcal{X}_{a_1}(\mathbf{x}_2) & \mathcal{X}_{a_2}(\mathbf{x}_2) & \dots & \mathcal{X}_{a_K}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \mathcal{X}_{a_1}(\mathbf{x}_N) & \mathcal{X}_{a_2}(\mathbf{x}_N) & \dots & \mathcal{X}_{a_K}(\mathbf{x}_N) \end{bmatrix}.$$

Prior and Posterior Probability

Prior Probability

The corresponding conjugate prior is given by a Gaussian distribution

$$p(\mathbf{u}) = \mathcal{N}(\mathbf{u} | \mathbf{m}_0, \mathbf{S}_0).$$

Posterior Probability

Then the posterior probability is given by

$$p(\mathbf{u} | \mathbf{X}, \mathbf{t}, \sigma^2) = \mathcal{N}(\mathbf{u} | \mathbf{m}_N, \mathbf{S}_N),$$

where

$$\mathbf{m}_N = \mathbf{S}_N \left(\frac{1}{\sigma^2} (\mathbf{M}^{-1})^* \Phi^* \mathbf{t} + \mathbf{S}_0^{-1} \mathbf{m}_0 \right),$$

$$\mathbf{S}_N^{-1} = \frac{1}{\sigma^2} (\mathbf{M}^{-1})^* \Phi^* \Phi \mathbf{M}^{-1} + \mathbf{S}_0^{-1},$$

and $*$ denotes the conjugate transpose (Hermitian transpose).

Except adding \mathbf{M} , the proof is the same as that of Bishop(2006).

There are many criteria for the optimal design (Chaloner & Verdinelli 1995). In this linear model,

Bayesian alphabetic optimality

- A-optimality: Minimize trace $\left[\left(\frac{1}{\sigma^2} (\mathbf{M}^{-1})^* \Phi^* \Phi \mathbf{M}^{-1} + \mathbf{S}_0^{-1} \right)^{-1} \right]$
- D-optimality: Maximize det $\left[\frac{1}{\sigma^2} (\mathbf{M}^{-1})^* \Phi^* \Phi \mathbf{M}^{-1} + \mathbf{S}_0^{-1} \right]$
- et cetera

In this work,

- we focus on a subclass of designs, which is limited by orthonormal systems.

What is the designs which satisfy $\frac{1}{N} \Phi^* \Phi = \mathbf{I}$?

→ Orthogonal designs (Orthogonal arrays), Hedayat et al. (1999)

Posterior variance for orthogonal designs

$$\mathbf{S}_N = \left[\left(\frac{N}{\sigma^2} (\mathbf{M}^{-1})^* \mathbf{M}^{-1} + \mathbf{S}_0^{-1} \right)^{-1} \right].$$

Advantages

- Easy to calculate the posterior variance
- the well-balanced design
- Not necessary to search for designs

Disadvantages

- Not depend on the prior distribution
→ If the influence of the prior distribution is strong, we should search for the optimal design.
- The number of experiments (inputs) N : restricted to q^k (q : the number of levels of Factor, k : integer)

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In this work,

- Focused on a subclass of designs, which satisfies that $\frac{1}{N}\Phi^*\Phi = \mathbf{I}$
→ orthogonal designs
- Easy to calculate the posterior variance
- Not depend on the prior distribution
→ If the influence of the prior distribution is strong, we should search for the optimal design.

Further works

- What is the condition satisfies the orthogonal design is optimal?
- Can we apply this kind of projection to other models?
- Can we use the orthogonal model to search for the optimal design?

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Appendix: Orthogonal design 1

In signal processing, the range of frequencies is expressed by the maximum frequency.

In experimental designs, however, it is often necessary to express the range of frequencies in greater detail.

- The range of frequencies can be expressed in greater detail by using a bounded set $A(\subseteq \{0, 1\}^n)$ instead of a maximum frequency.

Let A represent all factors that might influence the result of the experiment.

Example

$n = 3, A = \{000, 100, 010, 001, 110\}$.

- 100: factor 1
- 010: factor 2
- 001: factor 3
- 110: interaction of factor 1 and factor 2

The other interactions don't influence the result of the experiment.

Orthogonal design 2

- For $\mathbf{a} = (a_1, a_2, \dots, a_n)$, $\mathbf{a}' = (a'_1, a'_2, \dots, a'_n) \in \{0, 1\}^n$, define $\mathbf{a} + \mathbf{a}' = (a_1 + a'_1, a_2 + a'_2, \dots, a_n + a'_n)$, where $+$ is over $GF(2)$.
- For $\mathbf{a} \in \{0, 1\}^n$, define $v(\mathbf{a}) = \{i | a_i \neq 0, 1 \leq i \leq n\}$.
- ① For $A \subseteq \{0, 1\}^n$, define

$$H_A = \begin{bmatrix} h_{1,1} & h_{1,2} & \dots & h_{1,n} \\ h_{2,1} & h_{2,2} & \dots & h_{2,n} \\ \vdots & \vdots & \ddots & \vdots \\ h_{k,1} & h_{k,2} & \dots & h_{k,n} \end{bmatrix},$$

where $h_{i,j} \in GF(q)$ and H_A satisfies the following condition:

The set $\{\mathbf{h}_{\cdot j} | j \in v(\mathbf{a} + \mathbf{a}')\}$, where $\mathbf{h}_{\cdot j}$ is the j -th column of H_A , is linearly independent over $GF(q)$ for any $\mathbf{a}, \mathbf{a}' \in A$.

- ② An orthogonal design C for $A \subseteq \{0, 1\}^n$ is defined by

$$C = \{\mathbf{x} | \mathbf{x} = \mathbf{r}H_A, \mathbf{r} \in GF(q)^k\}.$$

Orthogonal design: Example

Example

$q = 2, n = 4, A = \{0000, 1000, 0100, 0010, 0001, 1100, 1010, 1001\}$.

- 1 By the definition, the sets containing the following must be linearly independent:

the 1st, 2nd, and 3rd columns of H_A ;

the 1st, 2nd, and 4th columns of H_A ; and

the 1st, 3rd and 4th columns of H_A .

Then, the corresponding matrix H_A is given by

$$H_A = \begin{bmatrix} 0 & 0 & 1 & 1 \\ 0 & 1 & 0 & 1 \\ 1 & 0 & 0 & 0 \end{bmatrix}.$$

- 2 An orthogonal design C for A is defined by

$$\begin{aligned} C &= \{\mathbf{x} \mid \mathbf{x} = \mathbf{r}H_A, \mathbf{r} \in \{0, 1\}^3\} \\ &= \{0000, 1000, 0101, 1101, 0011, 1011, 0110, 1110\}. \end{aligned}$$

Example: Posterior variance

Example: Posterior variance ($n=2$, $q=3$, $N=9$)

$$\mathbf{S}_9 = \left[\left(\frac{9}{\sigma^2} (\mathbf{M}^{-1})^* \mathbf{M}^{-1} + \mathbf{S}_0^{-1} \right)^{-1} \right],$$

where

$$(\mathbf{M}^{-1})^* \mathbf{M}^{-1} = \frac{1}{9} \begin{bmatrix} 9 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 6 & 3 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 3 & 6 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 6 & 3 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 3 & 6 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 4 & 2 & 2 & 1 \\ 0 & 0 & 0 & 0 & 0 & 2 & 4 & 1 & 2 \\ 0 & 0 & 0 & 0 & 0 & 2 & 1 & 4 & 2 \\ 0 & 0 & 0 & 0 & 0 & 1 & 2 & 2 & 4 \end{bmatrix}.$$