# Computing Bayes: Bayesian computation from 1763 to 2017!

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#### Bayes on the Beach, November, 2017

- The Royal Society, London, December 23, 1763:
- Richard Price read:

An Essay Towards Solving a Problem in the Doctrine of Chances by Reverend Thomas Bayes

- Three years after Bayes' death
- 'Bayesian inference' has its birth......

#### • The question posed?

- If perform *n* Bernoulli trials, with  $\theta = \text{probability of 'success'}$ 
  - Rolling a ball across a 'billiard' table n times
  - $\bullet \, \Rightarrow \,$  'success' if ball lands within a particular distance from the edge
- And record:  $\mathbf{y} = (1, 1, 0, 1, 0, ....0)'$

What is:

$$\Pr{ob(a < \theta < b | \mathbf{y})}?$$

• The answer offered?

$$\Pr{ob}(a < heta < b | \mathbf{y}) = \int_{a}^{b} p(\theta | \mathbf{y}) d heta$$

• where:

$$p( heta|\mathbf{y}) = rac{L( heta|\mathbf{y})p( heta)}{p(\mathbf{y})} = ext{posterior pdf}$$

with:

- First application (we think...) of 'inverse probability'
- Given a set of observations (y)
- Produced according to an assumed probability distribution
  - (Bernoulli here ....)
- Can we **invert** the problem to make a **probability statement** about the *unknown and unobservable θ*?
- $\equiv$  'Bayesian inference' in our modern language....

• Computational challenge??

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- **Closed-form** solution for  $p(\theta|\mathbf{y})...(\mathbf{beta} \text{ density})$
- However:

$$\Pr{ob(a < heta < b|\mathbf{y})} = \int\limits_{a}^{b} p( heta|\mathbf{y}) d heta = `incomplete beta function`$$

- does not have a closed form!
- (and was not yet numerically tabulated!)
- And the fact that Bayes could not find an accurate numerical solution
- Has been proposed as a possible reason for his not publishing the work! (Stigler (1983) 'The History of Statistics')
- → computational issues a feature of 'Bayesian inference' from its birth!!

#### Reverend Thomas Bayes: 1701-1761:



• Why is a Presbyterian clergyman in the mid-1700's playing around with billiard balls and mathematics??

#### Protestant Reformation: 1517+

- October 31st 1517: Castle Church, Wittenberg, Germany
- Martin Luther (a monk) nails to the door: 95 'theses' or 'objections' to the workings of the Roman Catholic Church
- And so begins (the most publicized) break from the established Church of Rome
- The Swiss follow: Ulrich Zwingli, John Calvin (mid-1500s....)
- All 'reformers' or protesters'....creating the new **Protestant movement**
- Stepping outside of the authority of the Pope
- Advocating a more personal connection with God
- Including ordinary people appointing their own pastors

#### Protestant Reformation: 1517+

- Across the English Channel?
- Tumultuous time....
- Henry the 8th/Mary 1st/Elizabeth 1st
- 'Protestants' (Church of England variety...) have ascendancy under Elizabeth
- Simultaneously, in Scotland, Calvin's brand of Protestantism spreads
- $\Rightarrow$  Presbyterians
- By Bayes time (1701-1761): 'Non-conformist' (e.g. Presbyterians) and Church of England clergy dotted throughout the British Isles
- $\Rightarrow$  Reverend Thomas Bayes preaching in Tunbridge Wells (England) 1734 +

### The Scottish Enlightenment (1700s/1800s)

- An 'easy' gig! (Bryson (2010) 'At Home: a Short History of Private Life'!!)
- The odd sermon on Sunday...
- A fair bit of spare time!
- Time to explore ideas
- 'Gentleman' scholars
- (Bayes had studied both theology **and mathematics** at the University of **Edinburgh**)
- Ideas; discovery; questioning; scientific experimentation valued in the time of the **Enlightenment**
- .....so what we see with Bayes all makes sense......

#### Pierre-Simon Laplace: 1749–1827

- But **Bayes** dies early
- Work eventually publicized by Price....but appears to have disappeared from view thereafter
- Then along comes **Pierre**.....



#### Pierre-Simon Laplace: 1749–1827

- Appears to have discovered 'Bayes Theorem' independently (1770 + )
- Applied method of **inverse probability** to several problems, with priors determined via more abstract reasoning
- Along the way introduced the Laplace (analytical) approximation to (Bayesian) integrals!
- $\Rightarrow$  first computational solution to intractable Bayesian problems!!
- The method of **inverse probability** remained dominant in the 1800s
  - (Feinberg (2006), 'When did Bayesian Inference become "Bayesian"')

- Somewhat usurped in the 1900s by ('frequentist') notions of:
  - Maximum likelihood estimation and associated 'sampling properties' (Fisher, 1922)
  - Hypothesis testing/p-values/confidence intervals (Neyman/Pearson, 1930+)
- Despite works on 'Bayesian inference' by:
  - De Finetti (1930, 1937)
  - Jeffreys (1939)
  - Savage (1954)
  - Lindley (1965, 1971)
  - Arnold Zellner (1971)

### State of Play in 'Bayesian Inference' in 1970s?

#### • Zellner, 1971: 'Bayesian Inference in Econometrics'

- Key aspects of coverage?
  - Gaussian (and associated) distributions dominate
  - natural conjugate priors
  - + non-informative (Jeffreys) priors
  - ullet  $\Rightarrow$  analytical solutions for **posterior moments**
  - ullet  $\Rightarrow$  analytical solutions for marginal posteriors
  - $\bullet\,\Rightarrow\,$  analytical solutions for marginal likelihoods
  - ullet  $\Rightarrow$  analytical solutions for **predictives**

### State of Play in 'Bayesian Inference' in 1970s?

- Some use of **low-dimensional (deterministic) numerical** integration
- (+ use of numerical tabulations of common integrals)
- Some use of analytical approximations
- No mention of simulation-based computation.....
- However.....

### State of Play in Bayesian Computation in 1980s?

- Assumed DGPs (models) are becoming more complex and high-dimensional; e.g:
  - full models of the economy
  - more complex time series (e.g. unit root/cointegration) models
  - latent variable (including state space) models
  - :
- Neither Bayes with deterministic numerical integration
- Nor Bayes with analytical approximations
- was viable as a general inferential method
- Plus, computers speeding up!
- Enter stage left: simulation-based computation......

#### What IS the computational challenge in Bayes?

• Virtually all quantities of interest in Bayesian statistics can be expressed as:

$$E(g(\theta)|\mathbf{y}) = \int_{\theta} g(\theta) p(\theta|\mathbf{y}) d\theta$$

• for some  $g(\theta)$ :

$$\begin{split} E(\boldsymbol{\theta}|\mathbf{y}) &= \int_{\boldsymbol{\theta}} \boldsymbol{\theta} p(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta} \\ p(\theta_1|\mathbf{y}) &= \int_{\boldsymbol{\theta}} p(\theta_1|\boldsymbol{\theta}_{-1},\mathbf{y}) p(\boldsymbol{\theta}_{-1}|\mathbf{y}) d\boldsymbol{\theta}_{-1} \\ \Pr ob(a < \boldsymbol{\theta} < b|\mathbf{y}) &= \int_{\boldsymbol{\theta}} \mathbf{I}_{(a < \boldsymbol{\theta} < b)} p(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta} \\ p(y_{T+1}|\mathbf{y}) &= \int_{\boldsymbol{\theta}} p(y_{T+1}|\boldsymbol{\theta},\mathbf{y}) p(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta} \\ \bullet \text{ all } \equiv E(g(\boldsymbol{\theta})|\mathbf{y}) \text{ for some } g(\boldsymbol{\theta}) \end{split}$$

### What IS the computational challenge in Bayes?

- i.e implementing Bayes is all about evaluating integrals!!!
- $\equiv E(g(\theta)|\mathbf{y})$  for some  $g(\theta)$
- Only when assuming simple models
  - and standard including natural conjugate priors
- will such integrals ( $\equiv$  expectations) be available in closed form!
- For most empirically realistic models
- The integrals need to be estimated in some way.....
- Three main options:

#### **Bayesian Numerical Methods**

1. Deterministic numerical integration methods:

$$\int_{\boldsymbol{\theta}} g(\boldsymbol{\theta}) p(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta} = \int_{\boldsymbol{\theta}_1 \boldsymbol{\theta}_2} \int_{\boldsymbol{\theta}_p} g(\boldsymbol{\theta}) p(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta} \approx \sum_{j=1}^{G} \sum_{j=1}^{G} \dots \sum_{j=1}^{G} \dots$$

- Computational burden =  $G^p$
- 'curse' of dimensionality
- $\Rightarrow$  no good in high-dimensional case!

- 2. Analytical approximation of the integrand:  $\Rightarrow$  closed-form integrals
  - 'Laplace' method
  - Integrated Nested Laplace (INLA) method
  - Variational Bayes
  - All feasible, but only ever produce approximate results

#### 3. Stochastic simulation (or sampling) methods

- With modern computing power: 'exact' solutions are attainable
- Plus: a very natural way of thinking about **the estimation of** an expectation
- $\Rightarrow$  the dominant approach in the literature.....

• Given:

$$E(g( heta)|\mathbf{y}) = \int_{m{ heta}} g(m{ heta}) p(m{ heta}|\mathbf{y}) dm{ heta}$$

for some  $g(\theta)$ 

- All simulation methods involve:
  - sampling from  $p(\theta|\mathbf{y})$
  - and using that sample to estimate  $E(g(\theta)|\mathbf{y})$
- From Statistics 101: we estimate a **population mean** with a **sample mean!!**
- So, at the end of the day we will (usually) do two simple things:

Overview

Construct a sample mean of some function of M posterior draws:

$$\overline{g(\boldsymbol{\theta})} = rac{1}{M} \sum_{j=1}^{M} g(\boldsymbol{\theta}^{(j)})$$

- (Legitimately) use frequentist concepts to:
  - Construct a standard error that measures the accuracy of  $\overline{g(\theta)}$  as an estimate of  $E(g(\theta)|\mathbf{y})$
  - WLLN  $\Rightarrow$  consistency of  $g(\theta)$  as an estimate of  $E(g(\theta)|\mathbf{y})$ (as  $M \to \infty$ )
  - CLT  $\Rightarrow$  asymptotic normality of  $\overline{g(\theta)}$  (as  $M \to \infty$ )
- The hard part? Getting the draws from  $p(\theta|\mathbf{y})!$

Independent sampling: Monte Carlo sampling

- An **independent** sample from  $p(\theta|\mathbf{y})$  is ideal: each new draw brings 'fresh' information about  $p(\theta|\mathbf{y})$ 
  - $\Rightarrow$  high accuracy  $\equiv$  small (simulation) standard error
- Monte Carlo sampling produces an independent sample from  $p(\theta|\mathbf{y})$  directly
  - Great when  $p(\theta|\mathbf{y})$  is of a standard form but  $E(g(\theta)|\mathbf{y})$  is not!
  - Think of Bayes and his beta probability!
- But complex model  $\Rightarrow$  complex  $L(\theta|\mathbf{y}) \Rightarrow p(\theta|\mathbf{y})$  non-standard
- $\Rightarrow$  for **realistic** models:

 $p(\theta|\mathbf{y})$  cannot be simulated from directly

• Enter importance sampling.....

Independent sampling: importance sampling

- Kloek and (Herman) van Dijk (1978)
- Dutch econometricians. Why?
- Back to the Protestant reformation!!
- **1568** the (mainly) Protestant Dutch threw off the their imperial overlord: the Catholic Spanish
- Struck out independently.....invented the powerful mercentile state
  - $\Rightarrow$  a strong tradition in economics/econometrics
  - ullet  $\Rightarrow$  Econometric Institute of the Erasmus University Rotterdam
  - Kloek and van Dijk
- But back to the integrals!!!

- Importance sampling: Simple idea!
- Say have  $q(\theta|\mathbf{y}) pprox p(\theta|\mathbf{y})$ , and from which we can sample
- Estimate  $E(g(\theta)|\mathbf{y})$  as

$$\overline{g(\boldsymbol{\theta})}^{lS} = \sum_{j=1}^{M} \left( g(\boldsymbol{\theta}^{(j)}) w(\boldsymbol{\theta}^{(j)}) \right) / \sum_{j=1}^{M} w(\boldsymbol{\theta}^{(j)})$$

• Using draws of  $\theta$  from the importance density  $q(\theta|\mathbf{y})$ 

• where: 
$$w(oldsymbol{ heta}^{(j)}) = p^*(oldsymbol{ heta}^{(j)}|oldsymbol{y})/q(oldsymbol{ heta}^{(j)}|oldsymbol{y})$$

• for some **kernel**  $p^*$  of p:

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$$p(\theta|\mathbf{y}) = c \times p^*(\theta|\mathbf{y}) \propto L(\theta|\mathbf{y}) \times p(\theta)$$

Need to be able to evaluate  $L(\theta|\mathbf{y}) \propto p(\mathbf{y}|\theta)$ 

Independent sampling: importance sampling

- Great!! Problem solved??
- As long as we can write down the assumed **DGP** we are in business?
- Ummm.....how to choose  $q(\theta|\mathbf{y})$  to 'match'  $p(\theta|\mathbf{y})$  when  $\theta$  is of high dimension???
- Light bulb moment!
- Why not break a **high**-dimensional problem down into a sequence of **lower**-dimensional problems??

- Geman and Geman (1984), Gelfand and Smith (1990)
- Simple (and **revolutionary**!) idea:
- Hard to sample from a (complex) joint posterior
- **Easier** to sample from (lower dimensional; simpler) **conditional** posteriors
- Why?
- Conditioning always makes life easier
- Something that is unknown is treated (temporarily....) as known
- + Low-dimensional problems easier to deal with

• E.g., say have 
$$oldsymbol{ heta}=(oldsymbol{ heta}_1,oldsymbol{ heta}_2)'$$

$$p(oldsymbol{ heta}|oldsymbol{y})=p(oldsymbol{ heta}_1,oldsymbol{ heta}_2|oldsymbol{y})$$

- Draw  $\theta_1$  and  $\theta_2$  iteratively from  $p(\theta_1|\theta_2, \mathbf{y})$  and  $p(\theta_2|\theta_1, \mathbf{y})$
- Under regularity  $\Rightarrow$  yields draws from the **joint**:  $p(\theta|\mathbf{y})$
- Cost??
- Drawing sequentially via the conditionals creates **dependence** in the sample
- $\Rightarrow$  a Markov chain with invariant distribution equal to  $p(\theta|\mathbf{y})$
- Gibbs an example of a Markov chain Monte Carlo (MCMC) algorithm

- ullet  $\Rightarrow$  Need to verify conditions for convergence to  $p(m{ heta}|\mathbf{y})$
- $\Rightarrow$  Need to monitor convergence (and 'burn-in') in practice.....
- $\bullet \Rightarrow \mathsf{Need}$  more draws to produce the same level of accuracy as an independent sample
- All that done though....once we have the draws we do the usual simple things with them
- (Standard error formulae simply reflect the dependence in the draws)
- Gibbs sampling a good starting point in many complicated models
- Exploits the simplicity that comes from conditioning

- Take, for e.g. a state space model
  - with 'static' parameters  $\theta_1$  and random parameters  $\theta_2$   $(\dim(\theta_2) \ge n!)$
- $p( heta_1, heta_2 | \mathbf{y})$  will not be amenable to analytical treatment
- But:
  - $p(\theta_1|\theta_2, \mathbf{y})$  is often simple (reflecting a **linear regression** structure)
  - $p(\theta_2|\theta_1, \mathbf{y})$  exploits filtering techniques
- Can also introduce **auxiliary** latent variables in order to produce simple conditionals
- ullet  $\Rightarrow$  integrated out via the Gibbs procedure.....
- ullet  $\Rightarrow$  draws on parameters of interest retained

- Introduced by Tanner and Wong (1987) as 'data augmentation'
- Note:

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to be standard enough to be simulated from directly

 $L(\boldsymbol{\theta}|\mathbf{y}) \propto p(\mathbf{y}|\boldsymbol{\theta})$  needs to be available

- Important: even when DGP is available
- Typically, not **all** conditionals are standard and hence **can** be drawn from!
- (e.g.  $p(\theta_2|\theta_1, \mathbf{y})$  in a non-linear state space model)
- Trick? Draw from it indirectly
- By inserting another MCMC chain within Gibbs:
- Metropolis-Hastings (MH)
  - Metropolis (1953) Los Alamos (US)....nuclear physicists....inventors of the atomic bomb.....
- Magic! Insertion produces a hybrid chain with p(θ|y) still the invariant distribution.....

#### Bayesian Simulation Methods Metropolis-Hastings (MH) (within Gibbs) sampling

- The thrust of **MH** within Gibbs (applied to  $p(\theta_2|\theta_1, \mathbf{y})$  say)
  - Draw from  $p(\theta_2|\theta_1, \mathbf{y})$  via a candidate  $q(\theta_2) \approx p(\theta_2|\theta_1, \mathbf{y})$
  - (Note the dimension reduction via Gibbs.....)
  - Accept candidate draw of θ<sub>2</sub> with a probability that depends on the ratio:

$$\frac{p^*(\boldsymbol{\theta}_2|\boldsymbol{\theta}_1, \mathbf{y})}{q(\boldsymbol{\theta}_2|\boldsymbol{\theta}_1, \mathbf{y})}$$

Need to be able to evaluate  $p^*(\theta_2|\theta_1, \mathbf{y}) \Leftrightarrow$ 

Need to be able to evaluate  $L(\theta|\mathbf{y}) \propto p(\mathbf{y}|\theta)$ 

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#### Bayesian Simulation Methods Pseudo-marginal MCMC

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In summary.....all methods so far:

Require the evaluation of:  $L(\theta|\mathbf{y}) \propto p(\mathbf{y}|\theta)!!$ 

- What if this is not the case? E.g. continuous time models with unknown transitions (p(y<sub>t</sub>|y<sub>t-1</sub>, θ)?
- Or, if dim( $\mathbf{y}$ ) is so large in dimension that evaluation of  $p(\mathbf{y}|\boldsymbol{\theta})$  (product of *n* terms....) is essentially infeasible??
- For some problems rescue comes via the magic of an unbiased estimate of p(y|θ)!
- and the use of so-called pseudo-marginal MCMC methods

#### Bayesian Simulation Methods Pseudo-marginal MCMC

• Say have some:

$$\widehat{p(\mathbf{y}|\boldsymbol{\theta})}$$

where:

$$E_{\mathbf{u}}\left[\widehat{p(\mathbf{y}|\boldsymbol{\theta})}\right] = p(\mathbf{y}|\boldsymbol{\theta})$$

- $\mathbf{u} =$ the auxiliary **random variables** underpinning the estimate
- and are intimately related to model-specific latent random variables
- Make this additional source of uncertainty explicit:

$$\widehat{p(\mathbf{y}|\boldsymbol{\theta})} = g(\mathbf{y}|\boldsymbol{\theta},\mathbf{u})$$

# Bayesian Simulation Methods

Pseudo-marginal MCMC

• Apply usual trick  $\Rightarrow$  augment the 'unknowns' of the problem with u:

$$g(\theta, \mathbf{u}|\mathbf{y}) \propto g(\mathbf{y}|\theta, \mathbf{u})g(\mathbf{u})p(\theta)$$

- ullet  $\Rightarrow$  apply MCMC to the augmented space  $(m{ heta},m{u})$
- $\Rightarrow$  produce **marginal** inferences about heta
- ('**pseudo**' due to the true likelihood not being used....)
- What do we get?
- Is  $p(\theta|\mathbf{y})$  the invariant distribution of an MCMC algorithm applied to  $(\theta, \mathbf{u})$ ?
- Yes! Due to the unbiasedness of likelihood estimate!
  - Beaumont, 2003, Andrieu and Roberts, 2009

### Bayesian Simulation Methods Pseudo-marginal MCMC

- Pseudo-marginal applied in a state space model?
  - $\Rightarrow$  particle filtering-based estimate of  $p(\mathbf{y}|\boldsymbol{\theta}_1)$
  - $\Rightarrow$  **u** = vector of uniforms driving the **particle filter**
  - $\Rightarrow$  Particle MCMC (PMCMC) Andrieu et al. 2010

Releases the burden of having to

evaluate <u>all</u> components of  $p(\mathbf{y}|\boldsymbol{\theta})$ 

- E.g. some filtering methods require only **simulation** from the *transition* densities
- But measurement densities still need to be evaluated (in the particle weights)

- Finally, pseudo-marginal MCMC has been applied specifically to reduce computational load associated with evaluating p(y|θ) when dim(y) is large
- 'Big data'
- Quiroz, Villani, Kohn and Tran 2017 subsample the data to produce an unbiased estimate of  $p(\mathbf{y}|\boldsymbol{\theta})$

 $\bullet \Rightarrow \widehat{p(\boldsymbol{\theta}|\mathbf{y})}$ 

# 'Exact' Bayesian Inference

- All done??
- Have access to multiple simulation-based methods: MC/IS/MCMC/PM-MCMC
- to produce  $p(\theta|\mathbf{y})$
- i.e. exact Bayesian inference (up to simulation error)
- But.....how to conduct posterior inference on heta when:
  - The DGP p(y|θ) is intractable in a way that precludes use of exact (including pseudo-marginal) methods?
  - Or the dimension of θ so large that exploration/marginalization infeasible via exact methods?
  - Or when the expertise to produce a finely-tuned efficient **exact** algorithm is not available?
- Can/must resort to approximate Bayesian inference

# 'Approximate' Bayesian Inference

- Goal then is to produce an approximation to  $p(\theta|\mathbf{y})$ :
  - (i) Approximate Bayesian computation (ABC)
  - (ii) Bayesian Synthetic likelihood
    - (i) and (ii) nested under 3. Simulation methods
  - (iii) Variational Bayes
  - (iv) Integrated nested Laplace (INLA)
    - (iii) and (iv) nested under 2. Analytical approximation methods

# (i) Approximate Bayesian Computation

- Aim is to produce **draws** from an **approximation** to  $p(\theta|\mathbf{y})$
- and use draws to estimate that approximation
- The simplest (accept/reject) form of the algorithm:
  - Simulate  $(\theta^i)$ , i = 1, 2, ..., N, from  $p(\theta)$
  - Simulate pseudo-data  $\mathbf{z}^{i}$ , i = 1, 2, ..., N, from  $p(\mathbf{z}|\boldsymbol{\theta}^{i})$
  - 3 Select  $(\theta^i)$  such that:

$$d\{\eta(\mathbf{y}), \eta(\mathbf{z}^i)\} \leq \varepsilon$$

- $\eta(.)$  is a (vector) summary statistic
- d{.} is a distance criterion
- $\bullet~$  the tolerance  $\varepsilon~$  is arbitrarily small
- (Recent reviews: Marin, Publo, Robert and Ryder, 2011; Sisson and Fan, 2011; Drovandi, 2017)

# (i) Approximate Bayesian Computation

#### Note:

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### Evaluation of $L(\theta|\mathbf{y}) \propto p(\mathbf{y}|\theta)$ is not required

Only simulation of  $p(\mathbf{y}|\boldsymbol{\theta})$  is required

- In practice:  $\eta(.)$  is never sufficient  $\Rightarrow$
- i.e.  $\eta(.)$  does not reproduce information content of **y**
- Selected draws (as  $\varepsilon \to 0$ ) estimate  $p(\theta|\eta(\mathbf{y}))$  (not  $p(\theta|\mathbf{y}))$
- Selection of  $\eta(.)$
- And hence, **proximity** of  $p(\theta|\eta(\mathbf{y}))$  to  $p(\theta|\mathbf{y})$  still an open and hot topic!
- More after tea!!!!

# (ii) Bayesian Synthetic Likelihood

• ABC attempts to estimate  $p( heta|\eta(\mathbf{y}))$  via simulation

• Given:

 $p(\boldsymbol{\theta}|\boldsymbol{\eta}(\mathbf{y})) \propto p(\boldsymbol{\eta}(\mathbf{y})|\boldsymbol{\theta}) p(\boldsymbol{\theta})$ 

• in essence ABC approximates  $p(\eta(\mathbf{y})|\boldsymbol{\theta})$  via simulation, as:

$$p(\boldsymbol{\eta}(\mathbf{y})|\boldsymbol{\theta}) \approx rac{1}{N} \sum_{i=1}^{N} \mathcal{I}\left(d\{\boldsymbol{\eta}(\mathbf{y}), \boldsymbol{\eta}(\mathbf{z}^{i})\} \leq \varepsilon\right)$$

for the accept/reject version

- BSL (Price, Drovandi, Lee and Nott, 2017 ):
- Approximate  $p(\pmb{\eta}(\mathbf{y})|\pmb{\theta})$  as:

 $p_{S}(\boldsymbol{\eta}(\mathbf{y})|\boldsymbol{\theta}) \approx N(\mu_{N}(\boldsymbol{\theta}), \boldsymbol{\Sigma}_{N}(\boldsymbol{\theta}))$ 

where  $\mu_N(\theta)$  and  $\Sigma_N(\theta)$  are computed from N simulated draws of  $\eta(\mathbf{z})$  from  $p(\mathbf{z}|\theta^i)$ , for a given  $\theta$ 

- Draws from  $p_S(\theta|\eta(\mathbf{y}))$  are obtained by embedding  $p_S(\eta(\mathbf{y})|\theta)$  within (say) an MCMC algorithm
- Both ABC and BSL can thus be seen as versions of pseudo-marginal methods!
  - although inference is only ever conditional on  $\eta(\mathbf{y})$  (not  $\mathbf{y}$ )
  - and hence is only ever approximate.....
- Note, again:

Only simulation of  $p(\mathbf{y}|\boldsymbol{\theta})$  is required

# (iii) Variational Bayes

- Simultaneous with the development of new (simulation-based) approximation methods by statisticians/econometricians
- **Computer science/machine learning** community have been developing their own (deterministic) approximation tool:
- Variational inference/Variational Bayes
- In the spirit of calculus of variations  $\Rightarrow$
- Approximate  $p(\theta|\mathbf{y})$  by some  $q^*(\theta) \in Q$  s.t:

$$q^{*}(\boldsymbol{\theta}) = \underset{q(\boldsymbol{\theta}) \in Q}{\arg\min} KL\left(q(\boldsymbol{\theta}) || p(\boldsymbol{\theta} | \mathbf{y})\right) = E_{q(\boldsymbol{\theta})}\left[\log\left(\frac{p(\boldsymbol{\theta} | \mathbf{y})}{q(\boldsymbol{\theta})}\right)\right]$$

• Nice reviews by Ormerod and Wand, 2010 and Blei, Kucukelbir and McAuliffe, 2017

- Approximating  $p(\boldsymbol{\theta}|\mathbf{y})$  via simulation replaced by
- Approximating  $p(\boldsymbol{\theta}|\mathbf{y})$  via **optimization**
- Trade-off between:
  - Choosing q to be flexible enough to capture features of  $p(\theta|\mathbf{y})$
  - Choosing *q* to be tractable enough to enable efficient optimization
- **Problem?** If don't know  $p(\theta|\mathbf{y})$  how can we approximate it via:

$$q^*(oldsymbol{ heta}) = rgmin_{q(oldsymbol{ heta}) \in Q} {
m Argmin} KL\left(q(oldsymbol{ heta}) || p(oldsymbol{ heta}| {f y}) 
ight)???$$

•  $\Rightarrow$  minimizing *KL*  $\equiv$  maximizing:

$$E_{q(\boldsymbol{ heta})}\left[\log\left(rac{p(\mathbf{y}, \boldsymbol{ heta})}{q(\boldsymbol{ heta})}
ight)
ight]$$

• where  $p(\mathbf{y}, \boldsymbol{\theta}) = p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})$  is (assumed to be) available!

# (iii) Variational Bayes

• Further:

$$E_{q(\theta)}\left[\log\left(rac{p(\mathbf{y}, \mathbf{ heta})}{q(\mathbf{ heta})}
ight)
ight] \leq \log p(\mathbf{y}) = \int_{\mathbf{ heta}} p(\mathbf{y}|\mathbf{ heta}) p(\mathbf{ heta}) d\mathbf{ heta}$$

 $\bullet$   $\Rightarrow$  a lower bound for the marginal likelihood (or 'evidence')

- used as an approximation to  $p(\mathbf{y})$ 
  - which would typically be approximated as an additional step in simulation (e.g. MCMC) settings
- Critically: to implement VB:

Evaluation of  $p(\mathbf{y}|\boldsymbol{\theta})$  is required!

# (iii) Variational Bayes

- Note however!
- Barthelme and Chopin, 2014, 'Expectation-Propagation for Likelihood-Free Inference'
  - Use of VB principles to implement ABC
- Tran, Nott and Kohn, 2106, 'Variational Bayes with Intractable Likelihood'
  - Use an **unbiased estimate** of  $p(\mathbf{y}|\boldsymbol{\theta})$  within **VB**
- Ong, Nott, Tran, Sisson and Drovandi, 2106, 'Variational Bayes with Synthetic Likelihood'
  - Use a synthetic likelihood estimate of  $p(\mathbf{y}|\boldsymbol{\theta})$  within **VB**
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All loosen the requirements on the tractability of  $p(\mathbf{y}|\boldsymbol{\theta})$ 

### (iv) INLA Remember Pierre?

• Yet *another* stream of **approximate inference** builds on **Pierre's** simple idea for approximating an integral:

$$\int_{X} e^{\{nf(x)\}} dx \approx e^{\{nf(\widehat{x})\}} \int_{X} e^{\left\{\frac{-n|f''(\widehat{x})|}{2}(x-\widehat{x})^{2}\right\}} dx$$
$$= e^{\{nf(\widehat{x})\}} \sqrt{\frac{2\pi}{n|f''(\widehat{x})|}}$$

- **Optimization** needed to obtain  $\hat{x}$
- Building on Laplace (1774) and Tierney and Kadane, 1986
- Rue, Martino and Chopin, 2009
- apply this idea to a very broad class of models:
- 'latent Gaussian models' (or 'Gaussian process models')

- $\Rightarrow$  Integrated Nested Laplace (INLA) approx. of  $p(\theta|\mathbf{y})$ 
  - A combination of (nested) Laplace (LA) approximations
  - $\bullet\,$  Plus a (low-dimensional) numerical integration (IN) step
- Again: INLA (like VB) amounts to replacing simulation by optimization
  - $\Rightarrow$  much attention given to the matter of **numerical opt.** in the given model class
  - The **optimization** in **INLA** being over a high dimensional vector of latent states....
- Critically, the application of **INLA**:

#### Requires the evaluation of $p(\mathbf{y}|\boldsymbol{\theta})!$

• (Augmentation with other methods for dealing with the case where  $p(\mathbf{y}|\boldsymbol{\theta})$  is **intractable** is surely possible......)

## The 21st Century and Beyond?

- So.....where are we heading now?
- What does this wealth of computational developments mean: for the future of statistical inference?
- Back in 2008 I had just finished reading: 'The Story of French'
- An historical perspective on the language and its place in the world
- Coincidentally, I was asked to name and chair a debate between Christian Robert (Bayesian) and Russell Davidson (frequentist), entitled:

#### The 21st Century Belongs to Bayes

• Certain analogies between language and statistical paradigm became clear!

# The 21st Century and Beyond?

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### French = Linga franca until 20th century

- = characterized by clear, coherent rules of grammar
- = characterized by a strong sense of correct usage
- = Bayesian inferential paradigm!

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- English = Linga franca in 20th century +
  - = evolved quite differently
  - = freely borrowing from many other languages
  - = an amalgam of different approaches and structures
  - = Classical/frequentist inferential paradigm!

- According to the last chapter in The Story of French',
- the authors bravely assert that **in the 21st century** the elegant, logical and coherent language of French may regain its preeminence!
- Is it the same with **Bayes** ???
- In particular now armed as it is with this immense array of new computational tools!

# The 21st Century and Beyond?

- So elegance and coherence in approach:
- Quantifying uncertainty about what is **unknown** conditional on what is **known** using the language of probability:  $p(\theta|\mathbf{y})$
- Underpinned by the ability to compute  $p(\theta|\mathbf{y})$
- Whether 'exactly' or in some 'approximate' fashion
- in almost every imaginable situation.....
- Surely, our man in 1700's England with the billiard balls and the time to explore ideas.....
- Is now our man for the 21st century and beyond.....